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## RECURRENCE RATES AND HITTING-TIME DISTRIBUTIONS FOR RANDOM WALKS ON THE LINE<sup>1</sup>

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We consider random walks on the line given by a sequence of independent identically distributed jumps belonging to the strict domain of attraction of a stable distribution, and first determine the almost sure exponential divergence rate, as  $\varepsilon \rightarrow 0$ , of the return time to  $(-\varepsilon, \varepsilon)$ . We then refine this result by establishing a limit theorem for the hitting-time distributions of  $(x - \varepsilon, x + \varepsilon)$  with arbitrary  $x \in \mathbb{R}$ .

**1. Introduction and results.** We consider a recurrent random walk on  $\mathbb{R}$ ,  $S_0 := 0$  and  $S_n := X_1 + \cdots + X_n$ ,  $n \geq 1$ , where the  $X_i$  are i.i.d. random variables on  $(\Omega, \mathcal{F}, \mathbb{P})$  such that  $\frac{S_n}{A_n}$  converges, for positive real numbers  $A_n$ , in distribution to a stable random variable  $X$  with index  $\alpha$ . Necessarily (due to recurrence),  $\alpha \in [1, 2]$ , and the sequence  $(A_n)_{n \geq 1}$  is regularly varying of index  $\frac{1}{\alpha}$ , satisfying  $\sum_{n \geq 1} \frac{1}{A_n} = \infty$ .

To capture the speed at which recurrence appears, it is possible to specify, for such a walk, some deterministic sequences  $(\varepsilon_n)$  such that  $S_n \in (-\varepsilon_n, \varepsilon_n)$  infinitely often, or  $S_n \notin (-\varepsilon_n, \varepsilon_n)$  eventually, almost surely. This classical question was addressed, for example, in [5] and [3], the results of which have recently been extended in [4].

Here, we are going to study the number of steps it takes to return to some small neighborhood of the origin (or to hit a different small interval for the first time). For related work on random walks in the plane, intimately related to the  $\alpha = 1$  case of the present paper, we refer to [8].

As an additional standing assumption on our walk, we will always require the distribution of the jumps  $X_i$  to satisfy the Cramér condition

$$(1) \quad \limsup_{|t| \rightarrow \infty} |\mathbb{E}[e^{itX_1}]| < 1.$$

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This readily implies, in particular, that the event  $\Omega^* := \{S_n \neq 0 \forall n \geq 1\}$  has positive probability, and  $\Omega^*$  has probability one if and only if no individual path returning to the origin has positive probability.

As a warm-up we first determine the a.s. rate at which the variables

$$\mathbf{T}_\varepsilon := \min\{n \geq 1 : |S_n| < \varepsilon\}, \quad \varepsilon > 0,$$

diverge on  $\Omega^*$  as  $\varepsilon \rightarrow 0$ . Let  $\beta \in [2, \infty]$  be the exponent conjugate to  $\alpha$ , that is,  $\alpha^{-1} + \beta^{-1} = 1$ .

**THEOREM 1.** *In the present setup,*

$$(2) \quad \lim_{\varepsilon \rightarrow 0} \frac{\log \mathbf{T}_\varepsilon}{\log \varepsilon} = -\beta \quad \text{a.s. on } \Omega^*.$$

Our main objective then is to determine the precise order of magnitude and to study the asymptotic distributional behavior, as  $\varepsilon \rightarrow 0$ , of the more general hitting times of  $\varepsilon$ -neighborhoods of arbitrary given points  $x$  on the line. We shall, in fact, do so for the walk  $S'_n := S'_0 + S_n$ ,  $n \geq 0$ , with random initial position  $S'_0$ , independent of  $(S_n)_{n \geq 0}$  and having an arbitrary fixed distribution  $P$  on  $\mathbb{R}$ . For any  $x \in \mathbb{R}$  we thus let

$$\mathbf{T}_\varepsilon^x := \inf\{m \geq 1 : |S'_m - x| < \varepsilon\}$$

and  $\Omega_x^* := \{S'_n \neq x \forall n \geq 1\}$ . Outside  $\Omega_x^*$  we clearly have  $\lim_{\varepsilon \rightarrow 0} \mathbf{T}_\varepsilon^x = \min\{m \geq 1 : S'_m = x\}$ .

It is convenient to state the results in terms of, and work with, the strictly increasing continuous function  $G: [0, +\infty) \rightarrow [0, +\infty)$  with  $G(0) = 0$  which affinely interpolates the values  $G(n) = \sum_{k=1}^n \frac{1}{A_k}$ ,  $n \geq 1$ . We denote by  $G^{-1}$  its inverse function. Evidently,  $G(n) = o(n)$ . Moreover, by the direct half of Karamata's theorem (cf. Propositions 1.5.8 and 1.5.9a of [1]),  $G$  is regularly varying with index  $\frac{1}{\beta}$ , and satisfies

$$(3) \quad \frac{n}{A_n} = o(G(n)) \quad \text{if } \alpha = 1 \quad \text{while } \frac{n}{A_n} \sim \frac{G(n)}{\beta} \text{ in case } \alpha \in (1, 2].$$

We establish a result on convergence in distribution for  $\varepsilon G(\mathbf{T}_\varepsilon^x)$  conditioned on  $\Omega_x^*$  [while  $\varepsilon G(\mathbf{T}_\varepsilon^x) \rightarrow 0$  outside this set]. In the case  $\alpha = 1$ , the limit distribution is the same as for square integrable random walk on the plane; cf. [8]. Recall that  $X$  has a density  $f_X$ . For simplicity we set  $\gamma := 2f_X(0)\mathbb{P}(\Omega^*)$ .

**THEOREM 2.** *Assume that  $\alpha = 1$ , and fix any  $x \in \mathbb{R}$ . Conditioned on  $\Omega_x^*$ , the variables  $\varepsilon G(\mathbf{T}_\varepsilon^x)$  converge in law,*

$$\lim_{\varepsilon \rightarrow 0} \mathbb{P}(\gamma \varepsilon G(\mathbf{T}_\varepsilon^x) \leq t | \Omega_x^*) = \frac{t}{1+t} \quad \forall t > 0.$$

For  $\alpha \in (1, 2]$ , different limit distributions arise, and we obtain convergence in law of  $\mathbf{T}_\varepsilon^x$  to the  $\frac{1}{\beta}$ -stable subordinator at an independent exponential time:

**THEOREM 3.** *Assume that  $\alpha \in (1, 2]$ , and fix any  $x \in \mathbb{R}$ . Conditioned on  $\Omega_x^*$ , the variables  $\varepsilon G(\mathbf{T}_\varepsilon^x)$  converge in law,*

$$\lim_{\varepsilon \rightarrow 0} \mathbb{P} \left( \Gamma \left( \frac{1}{\beta} \right) \frac{\gamma}{\beta} \varepsilon G(\mathbf{T}_\varepsilon^x) \leq t \mid \Omega_x^* \right) = \Pr(\mathcal{E} \mathcal{G}_{1/\beta}^{1/\beta} \leq t) \quad \forall t > 0$$

or, equivalently,

$$\lim_{\varepsilon \rightarrow 0} \mathbb{P} \left( \left( \Gamma \left( \frac{1}{\beta} \right) \frac{\gamma}{\beta} \right)^\beta \frac{\mathbf{T}_\varepsilon^x}{G^{-1}(1/\varepsilon)} \leq t \mid \Omega_x^* \right) = \Pr(\mathcal{E}^\beta \mathcal{G}_{1/\beta} \leq t) \quad \forall t > 0,$$

where  $\mathcal{E}$  and  $\mathcal{G}_{1/\beta}$  are independent random variables,  $\Pr(\mathcal{E} > t) = e^{-t}$  and  $\mathcal{G}_{1/\beta}$  having the one-sided stable law of index  $\frac{1}{\beta}$  with Laplace transform  $\mathbb{E}[e^{-s\mathcal{G}_{1/\beta}}] = e^{-s^{1/\beta}}$ ,  $s > 0$ .

In particular, we have:

**COROLLARY 1.** *If  $(X_n)_{n \geq 1}$  is an i.i.d. sequence of centered random variables with variance 1, satisfying the Cramér condition, and  $x \in \mathbb{R}$ , then*

$$\lim_{\varepsilon \rightarrow 0} \mathbb{P}(2\mathbb{P}(\Omega^*) \varepsilon \sqrt{\mathbf{T}_\varepsilon^x} \leq t \mid \Omega_x^*) = \Pr \left( \frac{\mathcal{E}}{|\mathcal{N}|} \leq t \right) \quad \forall t > 0$$

or, equivalently,

$$\lim_{\varepsilon \rightarrow 0} \mathbb{P}(4\mathbb{P}(\Omega^*)^2 \varepsilon^2 \mathbf{T}_\varepsilon^x \leq t \mid \Omega_x^*) = \Pr \left( \left( \frac{\mathcal{E}}{|\mathcal{N}|} \right)^2 \leq t \right) \quad \forall t > 0,$$

where  $\mathcal{E}$  and  $\mathcal{N}$  are independent variables,  $\mathcal{N}$  having a standard Gaussian distribution  $\mathcal{N}(0, 1)$ .

As Cheliotis does in [4], we will use the following extension of Stone's local limit theorem [9].

**PROPOSITION 1.** *Let  $\theta$  be such that  $\limsup_{|t| \rightarrow \infty} |\mathbb{E}[e^{itX_1}]| < \theta < 1$ , and let  $c > 1$ . Then there exists a real number  $h_0 > 0$  and an integer  $n_0 \geq 1$  such that, for any  $n \geq n_0$ , for any interval  $I$  contained in  $[-h_0, h_0]$ , of length larger than  $\theta^n$ , we have*

$$c^{-1} f_X(0) |I| < \mathbb{P} \left( \frac{S_n}{A_n} \in I \right) < c f_X(0) |I|.$$

## 2. Almost sure convergence: Proof of Theorem 1.

**PROOF OF THEOREM 1.** To begin with, choose  $\theta$ ,  $c$  and  $h_0$  as in Proposition 1.

To first establish an estimate from below, we fix any  $\xi > 1$  and set  $\varepsilon_n := G(n)^{-\xi}$ . This makes the series  $\sum_n \mathbb{P}(|S_n| < \varepsilon_n)$  summable: Indeed, by regular variation and (3), we have  $\frac{\varepsilon_n}{A_n} > \theta^n$  for  $n$  large, while

$$\frac{\varepsilon_n}{A_n} = O\left(\frac{G(n) - G(n-1)}{G(n-1)^\xi}\right) = O\left(\int_{n-1}^n \frac{G'(t)}{G(t)^\xi} dt\right),$$

which is summable since  $\int_1^\infty \frac{G'(t)}{G(t)^\xi} dt = \left[\frac{G(t)^{1-\xi}}{1-\xi}\right]_1^\infty < \infty$ . In particular,  $(\frac{-\varepsilon_n}{A_n}, \frac{\varepsilon_n}{A_n}) \subseteq [-h_0, h_0]$  for large  $n$ . Proposition 1 therefore applies to these intervals and shows that  $\mathbb{P}(|S_n| < \varepsilon_n) = O(\frac{\varepsilon_n}{A_n})$  is summable as well. Hence, by the Borel–Cantelli lemma,  $\mathbb{P}(|S_n| < \varepsilon_n \text{ i.o.}) = 0$ . Since  $\varepsilon_n \searrow 0$ , we can conclude that  $\mathbf{T}_{\varepsilon_n} > n$  eventually, almost surely on  $\Omega^*$ , and we get  $\liminf_{n \rightarrow \infty} \frac{\log G(\mathbf{T}_{\varepsilon_n})}{-\log \varepsilon_n} \geq \frac{1}{\xi}$  a.s. on  $\Omega^*$ . Using monotonicity of  $\log G(\mathbf{T}_\varepsilon)$  and the fact that  $\varepsilon_{n+1} \sim \varepsilon_n$ , this extends from the  $\varepsilon_n$  to the full limit as  $\varepsilon \rightarrow 0$ , and since  $\xi > 1$  was arbitrary, we conclude that

$$(4) \quad \liminf_{\varepsilon \rightarrow 0} \frac{\log G(\mathbf{T}_\varepsilon)}{-\log \varepsilon} \geq 1 \quad \text{a.s. on } \Omega^*.$$

To control the corresponding limsup, we now fix any  $\xi \in (0, 1)$ . From Proposition 1, using intervals  $(\frac{-\varepsilon_n}{A_n}, \frac{\varepsilon_n}{A_n})$  and regular variation of  $(A_n)_{n \geq 1}$ , we see that there exists a constant  $c' > 0$  such that for every  $\varepsilon \in (0, 1)$  there is some  $m_\varepsilon$  satisfying

$$\mathbb{P}(|S_k| < \varepsilon) \geq \frac{c' \varepsilon}{A_k} \quad \text{for } k \geq m_\varepsilon.$$

More precisely, the dependence of  $m_\varepsilon$  on  $\varepsilon$  comes from the requirement  $2\varepsilon/A_k > \theta^k$  for  $k \geq m_\varepsilon$  on the length of intervals, which is met by taking  $m_\varepsilon := \kappa(-\log \varepsilon)$  with a suitable constant  $\kappa > 0$ . Next, choose integers  $n_\varepsilon$  in such a way that  $G(n_\varepsilon) \leq \varepsilon^{-1/\xi} < G(n_\varepsilon + 1)$ . Inspired by a decomposition used by Dvoretzki and Erdős [6], we consider the pairwise disjoint events  $E_k^\varepsilon := \{|S_k| < \varepsilon \text{ and } \forall j = k+1, \dots, n_\varepsilon: |S_j - S_k| > 2\varepsilon\}$ ,  $1 \leq k \leq n_\varepsilon$ . By independence and stationarity we have

$$1 \geq \sum_{k=m_\varepsilon}^{n_\varepsilon} \mathbb{P}(E_k^\varepsilon) \geq \sum_{k=m_\varepsilon}^{n_\varepsilon} \mathbb{P}(|S_k| < \varepsilon) \mathbb{P}(\mathbf{T}_{2\varepsilon} > n_\varepsilon - k) \geq c' \varepsilon \mathbb{P}(\mathbf{T}_{2\varepsilon} > n_\varepsilon) \sum_{k=m_\varepsilon}^{n_\varepsilon} \frac{1}{A_k}.$$

Combining this with  $G(m_\varepsilon) = o(G(n_\varepsilon))$  [note that  $G(m_\varepsilon)$  is slowly varying], we obtain

$$\begin{aligned} \mathbb{P}(G(\mathbf{T}_{2\varepsilon}) > \varepsilon^{-1/\xi}) &\leq \mathbb{P}(G(\mathbf{T}_{2\varepsilon}) > G(n_\varepsilon)) = \mathbb{P}(\mathbf{T}_{2\varepsilon} > n_\varepsilon) \\ &\leq \frac{1}{c' \varepsilon (G(n_\varepsilon) - G(m_\varepsilon))} \sim \frac{\varepsilon^{1/\xi-1}}{c'}. \end{aligned}$$

Therefore, if we let  $\varepsilon_p := p^{-2/(1-\xi)}$ ,  $p \geq 1$ , the Borel–Cantelli lemma implies  $G(\mathbf{T}_{2\varepsilon_p}) \leq \varepsilon_p^{-1/\xi}$  eventually almost surely, showing that

$$\limsup_{p \rightarrow +\infty} \frac{\log G(\mathbf{T}_{2\varepsilon_p})}{-\log(2\varepsilon_p)} \leq \frac{1}{\xi}.$$

Using monotonicity as before, we can extend this from the  $\varepsilon_p$  to the full limit  $\varepsilon \rightarrow 0$ , and since this is true for any  $\xi \in (0, 1)$ , we obtain

$$(5) \quad \limsup_{\varepsilon \rightarrow 0} \frac{\log G(\mathbf{T}_\varepsilon)}{-\log(\varepsilon)} \leq 1 \quad \text{a.s. on } \Omega.$$

To conclude the proof, we note that for any  $\alpha \in [1, 2]$  we have

$$\lim_{n \rightarrow \infty} \frac{\log G(n)}{\log n} = \frac{1}{\beta},$$

which follows readily from regular variation of  $G$ ; compare Fact 2 in [4]. Together with (4) and (5), this entails

$$\lim_{\varepsilon \rightarrow 0} \frac{\log \mathbf{T}_\varepsilon}{-\log \varepsilon} = \lim_{\varepsilon \rightarrow 0} \frac{\log \mathbf{T}_\varepsilon}{\log G(\mathbf{T}_\varepsilon)} \cdot \frac{\log G(\mathbf{T}_\varepsilon)}{-\log \varepsilon} = \beta \quad \text{a.s. on } \Omega^*$$

as required.  $\square$

The first argument can easily be adapted to prove the lower bound (4) also for  $\mathbf{T}_\varepsilon^x$  with  $x \neq 0$ .

**3. Convergence in distribution for auxiliary processes.** We need to introduce auxiliary processes. Let  $(M_0^\varepsilon)_{\varepsilon > 0}$  be a family of random variables, independent of  $(S_n)_{n \geq 0}$ , such that  $M_0^\varepsilon$  has uniform distribution on the interval  $(-\varepsilon, \varepsilon)$ . For each  $\varepsilon > 0$  we define the walk  $(M_n^\varepsilon)_{n \geq 0}$  with random initial position  $M_0^\varepsilon$ , that is,  $M_n^\varepsilon := M_0^\varepsilon + S_n$ .

A major step toward Theorems 2 and 3 will be to prove a version which applies to the variables

$$\tau_\varepsilon := \min\{n \geq 1 : |M_n^\varepsilon| < \varepsilon\}, \quad \varepsilon > 0.$$

That is, we are interested in the limiting behavior, as  $\varepsilon \rightarrow 0$ , of the first return time distribution of the walk  $(M_n^\varepsilon)_{n \geq 0}$  to the interval  $(-\varepsilon, \varepsilon)$ . The goal of the present section is to establish:

**THEOREM 4.** *Assume that  $\alpha = 1$ . Conditioned on  $\Omega^*$ , the variables  $\varepsilon G(\tau_\varepsilon)$  converge in law,*

$$(6) \quad \lim_{\varepsilon \rightarrow 0} \mathbb{P}(\varepsilon G(\tau_\varepsilon) \leq t | \Omega^*) = \frac{t}{1+t} \quad \forall t > 0.$$

**THEOREM 5.** *Assume that  $\alpha \in (1, 2]$ . Conditioned on  $\Omega^*$ , the variables  $\varepsilon G(\tau_\varepsilon)$  converge in law,*

$$(7) \quad \lim_{\varepsilon \rightarrow 0} \mathbb{P} \left( \Gamma \left( \frac{1}{\beta} \right) \frac{\gamma}{\beta} \varepsilon G(\tau_\varepsilon) \leq t \mid \Omega^* \right) = \Pr(\mathcal{E} \mathcal{G}_{1/\beta} \leq t) \quad \forall t > 0.$$

*Equivalently,*

$$\lim_{\varepsilon \rightarrow 0} \mathbb{P} \left( \left( \Gamma \left( \frac{1}{\beta} \right) \frac{\gamma}{\beta} \right)^\beta \frac{\tau_\varepsilon}{G^{-1}(1/\varepsilon)} \leq t \mid \Omega^* \right) = \Pr(\mathcal{E}^\beta \mathcal{G}_{1/\beta} \leq t) \quad \forall t > 0.$$

Again we start with considerations valid for any  $\alpha \in [1, 2]$ . To begin with, we define, for  $\varepsilon > 0$ ,  $R > 0$ , and integers  $K > 0$ , auxiliary events

$$\Gamma_{\varepsilon, R, K} := \{\forall i = 1, \dots, K : S_i \neq 0 \text{ and } |M_i^\varepsilon| \leq R\},$$

which asymptotically exhaust  $\Omega^*$ , and on which we can work conveniently. As  $\varepsilon \rightarrow 0$  we have  $\mathbb{P}(\Gamma_{\varepsilon, R, K}) \rightarrow \mathbb{P}(\Gamma_{R, K})$  and  $\mathbb{P}(\Gamma_{\varepsilon, R, K} \setminus \Omega^*) \rightarrow \mathbb{P}(\Gamma_{R, K} \setminus \Omega^*)$ , where  $\Gamma_{R, K} := \{\forall i = 1, \dots, K : 0 < |S_i| \leq R\}$  (except, perhaps, for a countable set of  $R$ 's which we are going to avoid). Let  $n \in \mathbb{N}$ . Using again a decomposition similar to that of Dvoretzki and Erdős in [6], we find, for  $\varepsilon \in (0, \frac{1}{2})$ ,

$$(8) \quad \mathbb{P}(\Gamma_{\varepsilon, R, K}) = \sum_{k=0}^n p_k^- = \sum_{k=0}^n p_k^+$$

with  $p_k^\pm = p_{k, n, \varepsilon, R, K}^\pm := \mathbb{P}(\Gamma_{\varepsilon, R, K} \cap \{|M_k^\varepsilon| < \varepsilon \pm 2\varepsilon^2 \text{ and } \forall \ell = k+1, \dots, n : |M_\ell^\varepsilon| \geq \varepsilon \pm 2\varepsilon^2\})$  for  $1 \leq k \leq n$ , and  $p_0^\pm = p_{0, n, \varepsilon, R, K}^\pm := \mathbb{P}(\Gamma_{\varepsilon, R, K} \cap \{\forall \ell = 1, \dots, n : |M_\ell^\varepsilon| \geq \varepsilon \pm 2\varepsilon^2\})$ . In the sequel, we will use the following notation: given two functions  $a$  and  $b$ , the notation

$$a(\varepsilon, R, K) = o_{\varepsilon, R, K}(1) \quad \text{and} \quad b(R, K) = o_{R, K}(1)$$

will mean that

$$\limsup_{K \rightarrow +\infty} \limsup_{R \rightarrow +\infty} \limsup_{\varepsilon \rightarrow 0} |a(\varepsilon, R, K)| = 0 \quad \text{and} \quad \limsup_{K \rightarrow +\infty} \limsup_{R \rightarrow +\infty} |b(R, K)| = 0.$$

We will also write  $m_\varepsilon := (\log \varepsilon)^4$ . The following estimates are the basis of the argument to follow.

**LEMMA 1.** *Let  $c > 0$  and let, for every  $\varepsilon > 0$ ,  $n_\varepsilon$  be the integer such that  $G(n_\varepsilon) \leq \frac{c}{\varepsilon} < G(n_\varepsilon + 1)$ .*

*For arbitrary  $\gamma'$  and  $\gamma''$  such that  $0 < \gamma' < 2f_X(0) < \gamma''$ , we have*

$$\mathbb{P}(\Gamma_{\varepsilon, R, K}) \geq \mathbb{P}(\tau_\varepsilon > n_\varepsilon) + \mathbb{P}(\Gamma_{\varepsilon, R, K}) \gamma' \varepsilon \sum_{k=m_\varepsilon}^{n_\varepsilon} \frac{\mathbb{P}(\tau_\varepsilon > n_\varepsilon - k)}{A_k} + o_{\varepsilon, R, K}(1)$$

and

$$\mathbb{P}(\Gamma_{\varepsilon,R,K}) \leq \mathbb{P}(\tau_\varepsilon > n_\varepsilon) + \mathbb{P}(\Gamma_{\varepsilon,R,K})\gamma''\varepsilon \sum_{k=m_\varepsilon}^{n_\varepsilon} \frac{\mathbb{P}(\tau_\varepsilon > n_\varepsilon - k)}{A_k} + o_{\varepsilon,R,K}(1).$$

PROOF. For the course of this proof, we simplify notation by suppressing the parameters  $\varepsilon$ ,  $R$ , and  $K$  in  $m_\varepsilon$ ,  $n_\varepsilon$ ,  $M_i^\varepsilon$ , and  $\Gamma_{\varepsilon,R,K}$ . We will apply (8) with  $n = n_\varepsilon$ . Also, let  $\nu := \varepsilon^2$ .

(i) Starting with the  $k = 0$  term, we see that

$$p_0^- \geq \mathbb{P}(\Gamma \cap \{\forall \ell = 1, \dots, n : |M_\ell| \geq \varepsilon\}) \geq \mathbb{P}(\Gamma \cap \{\tau_\varepsilon > n\}).$$

We now consider the case where  $m \leq k \leq n$ . Let  $\mathcal{A} := (2\nu\mathbb{Z}) \cap (-\varepsilon + 3\nu, \varepsilon - 3\nu)$ . Notice that the sets  $Q_a := (a - \nu, a + \nu)$  with  $a \in \mathcal{A}$  are disjoint and contained in  $(-\varepsilon + 2\nu, \varepsilon - 2\nu)$ . Therefore the  $k$ th term in (8) satisfies

$$\begin{aligned} p_k^- &\geq \sum_{a \in \mathcal{A}} \mathbb{P}(\Gamma \cap \{M_k \in Q_a \text{ and } \forall \ell = k+1, \dots, n : |M_\ell| \geq \varepsilon - 2\nu\}) \\ (9) \quad &\geq \sum_{a \in \mathcal{A}} \mathbb{P}(\Gamma \cap \{M_k \in Q_a \text{ and } \forall \ell = k+1, \dots, n : |S_\ell - S_k + a| \geq \varepsilon - \nu\}) \\ &= \sum_{a \in \mathcal{A}} \mathbb{P}(\Gamma \cap \{M_k \in Q_a\}) \mathbb{P}(\forall \ell = 1, \dots, n-k : |S_\ell + a| \geq \varepsilon - \nu) \end{aligned}$$

by independence [where we assume that  $\varepsilon$  is so small that  $(\log \varepsilon)^4 > K$ ]. Note that

$$\begin{aligned} &\mathbb{P}(\Gamma \cap \{M_k \in Q_a\}) \\ &= \int_{\{\forall i : x_i \neq x_0, |x_i| \leq R\}} \mathbb{P}(S_{k-K} \in Q_a - x_K) d\mathbb{P}_{(M_0, \dots, M_K)}(x_0, \dots, x_K) \end{aligned}$$

with  $d\mathbb{P}_{(M_0, \dots, M_K)}$  denoting the distribution of  $(M_0, \dots, M_K)$ . Now fix  $\theta$  as in Proposition 1, and  $c \in (0, 1)$  such that  $\gamma' < 2f_X(0)/c$ . Elementary considerations show that Proposition 1 applies to  $I = \frac{1}{A_{k-K}}(Q_a - x_K)$  if  $\varepsilon$  is sufficiently small, and in this case gives

$$(10) \quad \mathbb{P}(\Gamma \cap \{M_k \in Q_a\}) \geq \mathbb{P}(\Gamma) \frac{\gamma'\nu}{A_k}.$$

Using this, plus the observation that conditioning on  $\{M_0 \in Q_a\}$  amounts to looking at  $M_n^* := M_0^* + S_n$ ,  $n \geq 0$ , with  $M_0^*$  uniformly distributed on  $Q_a$ , we can continue to estimate, for small  $\varepsilon$ ,

$$p_k^- \geq \mathbb{P}(\Gamma) \frac{\gamma'\nu}{A_k} \sum_{a \in \mathcal{A}} \mathbb{P}(\forall \ell = 1, \dots, n-k : |S_\ell + a| \geq \varepsilon - \nu)$$



$$\begin{aligned}
&\geq \mathbb{P}(\Gamma) \frac{\gamma' \nu}{A_k} \sum_{a \in \mathcal{A}} \mathbb{P}(\{\forall \ell = 1, \dots, n-k : |M_\ell| \geq \varepsilon\} | \{M_0 \in Q_a\}) \\
(11) \quad &\geq \mathbb{P}(\Gamma) \frac{\gamma' \varepsilon}{A_k} \sum_{a \in \mathcal{A}} \mathbb{P}(\{\forall \ell = 1, \dots, n-k : |M_\ell| \geq \varepsilon\} \cap \{M_0 \in Q_a\}) \\
&\geq \mathbb{P}(\Gamma) \frac{\gamma' \varepsilon}{A_k} (\mathbb{P}(\forall \ell = 1, \dots, n-k : |M_\ell| \geq \varepsilon) - \mathbb{P}(\varepsilon - 4\nu \leq |M_0| \leq \varepsilon)) \\
&= \mathbb{P}(\Gamma) \frac{\gamma' \varepsilon}{A_k} (\mathbb{P}(\tau_\varepsilon > n-k) - 8\nu).
\end{aligned}$$

Putting together these estimates via equation (8) gives

$$\begin{aligned}
&\mathbb{P}(\Gamma \cap \{\tau_\varepsilon > n\}) + \mathbb{P}(\Gamma) \gamma' \varepsilon \sum_{k=m}^n \frac{\mathbb{P}(\tau_\varepsilon > n-k)}{A_k} \\
&\leq \mathbb{P}(\Gamma) + \mathbb{P}(\Gamma) 8\gamma' \varepsilon \nu (G(n) - G(m)).
\end{aligned}$$

We observe that  $\Gamma^c \cap \{\tau_\varepsilon > n\} \subseteq \bigcup_{i=1}^K \{|M_i| > R\}$  for  $\varepsilon$  so small that  $n = n_\varepsilon > K$ . Since  $\limsup_{K \rightarrow +\infty} \limsup_{R \rightarrow +\infty} \limsup_{\varepsilon \rightarrow 0} \mathbb{P}(\bigcup_{i=1}^K \{|M_i| > R\}) = 0$  and  $\lim_{\varepsilon \rightarrow 0} \varepsilon^3 (G(n) - G(m)) = 0$ , this proves the first assertion of the lemma.

(ii) We only provide a sketch of the proof of the second point since the arguments are very similar to the above. Using (8) gives

$$\mathbb{P}(\Gamma) \leq \mathbb{P}(\Gamma \cap \{\tau_\varepsilon > n\}) + \mathbb{P}(\Gamma \setminus \Omega^*) + \mathbb{P}(\Omega^* \cap \{\tau_{3\varepsilon} \leq m\}) + \sum_{k=m}^n p_k^+$$

since  $\sum_{k=1}^m p_k^+ \leq \mathbb{P}(\Gamma \cap \{\tau_{3\varepsilon} \leq m\})$ . Next, take  $\bar{\mathcal{A}} := (2\nu\mathbb{Z}) \cap (-\varepsilon - 3\nu, \varepsilon + 3\nu)$  and intervals  $\bar{Q}_a := [a - \nu, a + \nu]$ ,  $a \in \bar{\mathcal{A}}$ , which cover  $(-\varepsilon - 2\nu, \varepsilon + 2\nu)$ . We can then use arguments parallel to those of part (i) to obtain

$$\begin{aligned}
\sum_{k=m}^n p_k^+ &\leq \sum_{k=m}^n \sum_{a \in \bar{\mathcal{A}}} \mathbb{P}(\Gamma \cap \{M_k \in \bar{Q}_a \text{ and } \forall \ell = k+1, \dots, n : |M_\ell| > \varepsilon + 2\nu\}) \\
&\vdots \\
&\leq \mathbb{P}(\Gamma) \gamma'' \varepsilon \sum_{k=m}^n \frac{\mathbb{P}(\tau_\varepsilon > n-k)}{A_k} + \mathbb{P}(\Gamma) 8\gamma'' \varepsilon \nu (G(n) - G(m)),
\end{aligned}$$

which proves our claim since  $\lim_{\varepsilon \rightarrow 0} \mathbb{P}(\Omega^* \cap \{\tau_{3\varepsilon} \leq m\}) = 0$  as a consequence of Theorem 1 and since  $\mathbb{P}(\Gamma \setminus \Omega^*) = o_{\varepsilon, R, K}(1)$ .  $\square$

This enables us to derive an asymptotic bound for the tails of the distributions of the  $\varepsilon G(\tau_\varepsilon)$  as  $\varepsilon \rightarrow 0$ .

LEMMA 2. For all  $\alpha \in [1, 2]$  and any  $t > 0$  we have

$$\limsup_{\varepsilon \rightarrow 0} \mathbb{P}(\gamma \varepsilon G(\tau_\varepsilon) > t) \leq \frac{\mathbb{P}(\Omega^*)}{1+t}.$$

PROOF. Fix  $t, R, K$  and  $0 < \gamma' < 2f_X(0)$ . For  $\varepsilon > 0$  choose  $n_\varepsilon$  so that  $G(n_\varepsilon) \leq \frac{t}{\gamma \varepsilon} \leq G(n_\varepsilon + 1)$ , whence  $\mathbb{P}(\varepsilon \gamma G(\tau_\varepsilon) > t) \sim \mathbb{P}(\tau_\varepsilon > n_\varepsilon)$ . Recall that  $m_\varepsilon := (\log \varepsilon)^4$ . As in the proof of Theorem 1 we see that  $G(m_\varepsilon) = o(G(n_\varepsilon))$ . Therefore

$$(12) \quad \varepsilon \sum_{k=m_\varepsilon}^{n_\varepsilon} \frac{\mathbb{P}(\tau_\varepsilon > n_\varepsilon - k)}{A_k} \geq \varepsilon(G(n_\varepsilon) - G(m_\varepsilon))\mathbb{P}(\tau_\varepsilon > n_\varepsilon) \sim \frac{t}{\gamma} \mathbb{P}(\tau_\varepsilon > n_\varepsilon).$$

Together with the first part of Lemma 1, this yields

$$\limsup_{\varepsilon \rightarrow 0} \mathbb{P}(\varepsilon \gamma G(\tau_\varepsilon) > t) \leq \frac{\mathbb{P}(\Gamma_{R,K}) + o_{R,K}(1)}{1 + (t\gamma'/\gamma)\mathbb{P}(\Gamma_{R,K})}.$$

Taking successively  $R \rightarrow \infty$ , then  $K \rightarrow \infty$  and finally  $\gamma' \rightarrow 2f_X(0)$ , we obtain the lemma.  $\square$

When  $\alpha = 1$ , this upper bound actually is the limit:

LEMMA 3. If  $\alpha = 1$ , then for any  $t > 0$  we have

$$\liminf_{\varepsilon \rightarrow 0} \mathbb{P}(\gamma \varepsilon G(\tau_\varepsilon) > t) \geq \frac{\mathbb{P}(\Omega^*)}{1+t}.$$

PROOF. Fix  $t, R, K$  and  $\gamma'' > 2f_X(0)$ , and choose  $m_\varepsilon$  and  $n_\varepsilon$  as in the previous proof.

Since  $\alpha = 1$  means that  $G$  is slowly varying, we have  $G(2n_\varepsilon) - G(n_\varepsilon) = o(G(n_\varepsilon))$ . Hence

$$(13) \quad \begin{aligned} & \mathbb{P}(\tau_\varepsilon > 2n_\varepsilon) + \mathbb{P}(\Gamma_{\varepsilon,R,K})\gamma''\varepsilon \sum_{k=m_\varepsilon}^{2n_\varepsilon} \frac{\mathbb{P}(\tau_\varepsilon > 2n_\varepsilon - k)}{A_k} \\ & \leq \mathbb{P}(\tau_\varepsilon > n_\varepsilon) + \mathbb{P}(\Gamma_{\varepsilon,R,K})\gamma''\varepsilon \left( \sum_{k=m_\varepsilon}^{n_\varepsilon} \frac{\mathbb{P}(\tau_\varepsilon > n_\varepsilon)}{A_k} + \sum_{k=n_\varepsilon}^{2n_\varepsilon} \frac{1}{A_k} \right) \\ & \leq \mathbb{P}(\tau_\varepsilon > n_\varepsilon) + \mathbb{P}(\Gamma_{\varepsilon,R,K})\gamma''\varepsilon G(n_\varepsilon)[\mathbb{P}(\tau_\varepsilon > n_\varepsilon) + o(1)] \\ & \leq \mathbb{P}(\tau_\varepsilon > n_\varepsilon) + t \frac{\gamma''}{\gamma} \mathbb{P}(\Gamma_{\varepsilon,R,K})\mathbb{P}(\tau_\varepsilon > n_\varepsilon) + o(1). \end{aligned}$$

Combining these observations with the second estimate of Lemma 1 (replacing  $n_\varepsilon$  by  $2n_\varepsilon$ ) entails

$$\liminf_{\varepsilon \rightarrow 0} \mathbb{P}(\tau_\varepsilon > n_\varepsilon) \geq \frac{\mathbb{P}(\Gamma_{R,K}) - o_{R,K}(1)}{1 + (t\gamma''/\gamma)\mathbb{P}(\Gamma_{R,K})}.$$

We conclude by successively taking  $R \rightarrow \infty$ ,  $K \rightarrow \infty$  and  $\gamma'' \rightarrow 2f_X(0)$ .  $\square$

PROOF OF THEOREM 4. Immediate from Lemmas 2 and 3, as  $\varepsilon G(\tau_\varepsilon) \rightarrow 0$  outside  $\Omega^*$ .  $\square$

When  $\alpha \in (1, 2]$ , Lemma 1 does not yet give the limit distribution. Still, it immediately implies the tightness of the family of distributions with the normalization given there:

LEMMA 4. *The family of distributions of the random variables  $\varepsilon G(\tau_\varepsilon)$ ,  $\varepsilon \in (0, 1)$ , is tight.*

Hence it will be enough to prove that the advertised limit law is the only possible accumulation point of our distributions. We henceforth abbreviate

$$Z_\varepsilon := \frac{\gamma}{\beta} \varepsilon G(\tau_\varepsilon), \quad \varepsilon > 0.$$

LEMMA 5. *Suppose that  $\alpha \in (1, 2]$ . Let  $(\varepsilon_p)_{p \geq 1}$  be a positive sequence with  $\lim_{p \rightarrow \infty} \varepsilon_p = 0$ , and such that the conditional distributions of the  $Z_{\varepsilon_p}$  on  $\Omega^*$  converge to the law of some random variable  $Y$ . Then its tail satisfies the integral equation*

$$1 = \Pr(Y > t) + t \int_0^1 \frac{\Pr(Y > t(1-u)^{1/\beta})}{u^{1/\alpha}} du \quad \forall t > 0.$$

PROOF. (i) We write  $f(t) := \Pr(Y > t)$ , and first prove that

$$\forall t > 0 \quad 1 \geq f(t) + t \int_0^1 u^{-1/\alpha} f(t(1-u)^{1/\beta}) du.$$

Let us only consider  $\varepsilon$  belonging to  $\{\varepsilon_p, p \geq 1\}$ . Note that by monotonicity and right continuity of  $f$  it suffices to prove the inequality for all  $t \in (0, \infty)$  such that, for all  $N \geq 1$  and all  $r = 0, \dots, N-1$ , the function  $f$  is continuous at  $t(1 - \frac{r}{N})^{1/\beta}$ . Henceforth such a  $t$  will be fixed.

Now take some  $\delta > 0$ . We claim that one can choose  $N_\delta > 1$  such that for all  $N \geq N_\delta$ ,

$$(14) \quad \left| \int_0^1 \frac{f(t(1-u)^{1/\beta})}{u^{1/\alpha}} du - \frac{1}{N} \sum_{r=1}^{N-1} \frac{f(t(1 - (r/N))^{1/\beta})}{((r+1)/N)^{1/\alpha}} \right| \leq \delta.$$

Indeed, take  $\Delta \in (0, 1)$  such that  $\beta \Delta^{1/\beta} < \delta/4$ . For any  $N$  we have

$$\frac{1}{N} \sum_{r=1}^{\lfloor \Delta N \rfloor} \left( \frac{r+1}{N} \right)^{-1/\alpha} \leq \int_0^\Delta u^{-1/\alpha} du = \beta \Delta^{1/\beta} < \delta/4.$$

Since  $f$  is bounded by one this implies that both the integral in (14) restricted to  $[0, \Delta]$  and the sum from  $r = 1$  to  $\lfloor \Delta N \rfloor$  are bounded by  $\delta/4$ . The

claim follows by taking  $N_\delta$  so large that the approximation of the Riemann integral on the interval  $[\Delta, 1]$  by the Riemann sum with step  $1/N$  has a precision at least  $\delta/2$ .

Now fix integers  $N \geq N_\delta$ ,  $K \geq 1$ , and some  $0 < \gamma' < 2f_X(0)$ . For  $\varepsilon > 0$  small enough take  $n_\varepsilon$  such that  $G(n_\varepsilon) \leq \frac{\beta t}{\gamma \varepsilon} < G(n_\varepsilon + 1)$  [and hence  $G(n_\varepsilon) \sim \frac{\beta t}{\gamma \varepsilon}$ ].

According to the first point of Lemma 1, since  $\frac{n_\varepsilon}{N} \geq m_\varepsilon$ , we have

$$\mathbb{P}(\Gamma_{\varepsilon, R, K}) \geq \mathbb{P}(Z_\varepsilon > t) + \mathbb{P}(\Gamma_{\varepsilon, R, K}) \gamma' \varepsilon \sum_{k=n_\varepsilon/N}^{n_\varepsilon} \frac{\mathbb{P}(\tau_\varepsilon > n_\varepsilon - k)}{A_k} + o_{\varepsilon, R, K}(1).$$

Due to our assumption on the  $Z_{\varepsilon_p}$  and  $t$ , we see that  $\mathbb{P}(Z_\varepsilon > t) \rightarrow \mathbb{P}(\Omega^*)f(t)$  as  $\varepsilon_p \rightarrow 0$ . Next, by monotonicity,

$$\begin{aligned} & \sum_{k=n_\varepsilon/N}^{n_\varepsilon} \frac{\mathbb{P}(\tau_\varepsilon > n_\varepsilon - k)}{A_k} \\ & \geq \sum_{r=1}^{N-1} \sum_{k=0}^{n_\varepsilon/N-1} \frac{\mathbb{P}(\tau_\varepsilon > n_\varepsilon - k - (rn_\varepsilon/N))}{A_{k+(rn_\varepsilon/N)}} \\ & \geq \sum_{r=1}^{N-1} \left( G\left(\frac{r+1}{N}n_\varepsilon\right) - G\left(\frac{r}{N}n_\varepsilon\right) \right) \mathbb{P}\left(\tau_\varepsilon > \left(1 - \frac{r}{N}\right)n_\varepsilon\right). \end{aligned}$$

By regular variation, the first term of the product is asymptotically equivalent to

$$G(n_\varepsilon) \left[ \left(\frac{r+1}{N}\right)^{1/\beta} - \left(\frac{r}{N}\right)^{1/\beta} \right] \geq \frac{G(n_\varepsilon)}{\beta N ((r+1)/N)^{1/\alpha}}$$

as  $\varepsilon_p \rightarrow 0$ . On the other hand, the second term is equal to

$$\mathbb{P}\left(Z_\varepsilon > \varepsilon \frac{\gamma}{\beta} G\left(\left(1 - \frac{r}{N}\right)n_\varepsilon\right)\right) \rightarrow \mathbb{P}(\Omega^*)f\left(t\left(1 - \frac{r}{N}\right)^{1/\beta}\right),$$

since  $G\left(\left(1 - \frac{r}{N}\right)n_\varepsilon\right) \sim \left(1 - \frac{r}{N}\right)^{1/\beta} G(n_\varepsilon)$ . As a consequence, we see that

$$\begin{aligned} (15) \quad & \liminf_{p \rightarrow \infty} \varepsilon_p \sum_{k=n_{\varepsilon_p}/N}^{n_{\varepsilon_p}} \frac{\mathbb{P}(\tau_{\varepsilon_p} > n_{\varepsilon_p} - k)}{A_k} \\ & \geq \mathbb{P}(\Omega^*) \frac{t}{\gamma} \frac{1}{N} \sum_{r=1}^{N-1} \frac{f(t(1-r/N)^{1-1/\alpha})}{((r+1)/N)^{1/\alpha}} \\ & \geq \mathbb{P}(\Omega^*) \frac{t}{\gamma} \left( \int_0^1 \frac{f(t(1-u)^{1-1/\alpha})}{u^{1/\alpha}} du - \delta \right). \end{aligned}$$

Combining all these asymptotic estimates and taking the limit  $\varepsilon_p \rightarrow 0$ , we end then up with

$$\begin{aligned} \mathbb{P}(\Gamma_{R,K}) &\geq \mathbb{P}(\Omega^*) \left[ f(t) + \frac{\mathbb{P}(\Gamma_{R,K})\gamma't}{\gamma} \left( \int_0^1 \frac{f(t(1-u)^{1-1/\alpha})}{u^{1/\alpha}} du - \delta \right) \right] \\ &\quad + o_{R,K}(1). \end{aligned}$$

Successively letting  $R \rightarrow \infty$ ,  $K \rightarrow \infty$ ,  $\gamma' \rightarrow 2f_X(0)$  and  $\delta \rightarrow 0$  we obtain the desired inequality.

(ii) The converse inequality is proved analogously, using the other half of Lemma 1 with the following adaptation: we have

$$\mathbb{P}(\Gamma_{\varepsilon,R,K}) \leq \mathbb{P}(Z_\varepsilon > t) + \mathbb{P}(\Gamma_{\varepsilon,R,K})\gamma''\varepsilon \sum_{k=m_\varepsilon}^{n_\varepsilon} \frac{\mathbb{P}(\tau_\varepsilon > n_\varepsilon - k)}{A_k} + o_{\varepsilon,R,K}(1).$$

Since,  $G(n_\varepsilon/N) \sim G(n_\varepsilon)N^{-1/\beta}$  as  $\varepsilon$  goes to 0, we have, for  $\varepsilon$  small enough,

$$\varepsilon \sum_{k=m_\varepsilon}^{n_\varepsilon/N} \frac{\mathbb{P}(\tau_\varepsilon > n_\varepsilon - k)}{A_k} \leq \varepsilon G\left(\frac{n_\varepsilon}{N}\right) \leq 2\varepsilon G(n_\varepsilon)N^{-1/\beta} \leq 2\frac{\beta t}{\gamma} N^{-1/\beta}$$

and so

$$\begin{aligned} \mathbb{P}(\Gamma_{\varepsilon,R,K}) &\leq \mathbb{P}(Z_\varepsilon > t) + \mathbb{P}(\Gamma_{\varepsilon,R,K})\gamma''\varepsilon \sum_{k=n_\varepsilon/N}^{n_\varepsilon} \frac{\mathbb{P}(\tau_\varepsilon > n_\varepsilon - k)}{A_k} + o_{\varepsilon,R,K}(1) \\ &\quad + 2\gamma''\frac{\beta t}{\gamma} N^{-1/\beta}. \quad \square \end{aligned}$$

Now let us identify the limit distribution satisfying the equality given by Lemma 5. To this end we consider the variables

$$Z'_\varepsilon := \left(\frac{\gamma}{\beta}\right)^\beta \frac{\tau_\varepsilon}{G^{-1}(1/\varepsilon)}, \quad \varepsilon > 0.$$

LEMMA 6. *The conditional distributions of the  $Z_{\varepsilon_p}$  converge to a random variable  $Y$  iff the conditional distributions of the  $Z'_{\varepsilon_p}$  converge to  $Y^\beta$ . The latter then satisfies*

$$1 = \Pr(Y^\beta > t) + \int_0^t \frac{\Pr(Y^\beta > t-v)}{v^{1/\alpha}} dv \quad \forall t > 0.$$

PROOF. The equivalence of the two conditional distributional convergence statements follows from regular variation of  $G^{-1}$ ; see, for example, Lemma 1 of [2]. Suppose that they hold. Then, according to Lemma 5, for

any  $t > 0$ , we have

$$1 = \Pr(Y^\beta > t) + t^{1/\beta} \int_0^1 \frac{\Pr(Y^\beta > t(1-u))}{u^{1/\alpha}} du,$$

and the conclusion follows by a change of variables,  $v = tu$ .  $\square$

LEMMA 7. *Let  $W$  be a random variable with values in  $[0, \infty)$  satisfying*

$$(16) \quad \Pr(W \leq t) = \int_0^t \frac{\Pr(W > t-v)}{v^{1/\alpha}} dv \quad \forall t > 0.$$

Then

$$\mathbb{E}[e^{-sW}] = \frac{1}{1 + c_\beta s^{1/\beta}} \quad \forall s > 0$$

with  $c_\beta := \Gamma(\frac{1}{\beta})^{-1}$ . In particular, the distribution of  $W$  coincides with that of  $c_\beta^\beta \mathcal{E}^\beta \mathcal{G}_{1/\beta}$ , where the independent variables  $\mathcal{E}$  and  $\mathcal{G}_{1/\beta}$  are as in the statement of Theorem 3.

PROOF. Let  $s > 0$ . We have

$$\begin{aligned} \mathbb{E}[e^{-sW}] &= \int_0^{+\infty} \Pr(e^{-sW} \geq u) du \\ &= \int_0^{+\infty} \Pr\left(W \leq -\frac{\log(u)}{s}\right) du \\ &= \int_0^{+\infty} \Pr(W \leq v) s e^{-sv} dv. \end{aligned}$$

Hence, for any  $s > 0$ , we find

$$\begin{aligned} \mathbb{E}[e^{-sW}] &= \int_0^{+\infty} \left[ \int_0^v \frac{\Pr(W \geq v-w)}{w^{1/\alpha}} dw \right] s e^{-sv} dv \\ &= \int_0^{+\infty} \frac{1}{w^{1/\alpha}} \left[ \int_w^{+\infty} \Pr(W \geq v-w) s e^{-sv} dv \right] dw \\ &= \int_0^{+\infty} \frac{e^{-sw}}{w^{1/\alpha}} \left[ \int_0^{+\infty} \Pr(W \geq z) s e^{-sz} dz \right] dw \\ &= \int_0^{+\infty} \frac{e^{-sw}}{w^{1/\alpha}} \left[ 1 - \int_0^{+\infty} \Pr(W \leq z) s e^{-sz} dz \right] dw \\ &= \int_0^{+\infty} \frac{e^{-sw}}{w^{1/\alpha}} dw \cdot [1 - \mathbb{E}[e^{-sW}]], \end{aligned}$$

and our claim about the Laplace transform of  $W$  follows since

$$\int_0^{+\infty} \frac{e^{-sw}}{w^{1/\alpha}} dw = \frac{\beta}{s^{1/\beta}} \int_0^{+\infty} e^{-z^\beta} dz = \frac{1}{c_\beta s^{1/\beta}} \quad \text{with } c_\beta := \frac{1}{\Gamma(1/\beta)}.$$

Given this, a routine calculation (cf. Problem XIII.11.10 of [7]) shows that  $W$  indeed has the same Laplace transform as  $c_\beta^\beta \mathcal{E}^\beta \mathcal{G}_{1/\beta}$ .  $\square$

**PROOF OF THEOREM 5.** According to Lemma 4 the family of distributions of the  $Z_\varepsilon$ ,  $\varepsilon \in (0, 1)$ , is tight. By Lemmas 5, 6 and 7, the law of  $c_\beta \mathcal{E} \mathcal{G}_{1/\beta}^{1/\beta}$  is the only possible accumulation point of these distributions.  $\square$

**4. Convergence in distribution for  $\mathbf{T}_\varepsilon^x$ .** To complete the proof of Theorems 2 and 3 we now utilize Theorems 4 and 5. Note first that it suffices to prove Theorems 2 and 3 under the additional assumption that  $S'_0 = 0$ , in which case

$$\mathbf{T}_\varepsilon^x = \hat{\mathbf{T}}_\varepsilon^x := \inf\{n \geq 1 : |S_n - x| < \varepsilon\} \quad \text{and} \quad \Omega_x^* = \hat{\Omega}_x^* := \{S_n \neq x \ \forall n\}.$$

Indeed, in the situation of Theorem 2, with arbitrary distribution  $P$  of  $S'_0$ , we then have

$$\mathbb{P}(\gamma \varepsilon G(\mathbf{T}_\varepsilon^x) \leq t) = \int_{\mathbb{R}} \mathbb{P}(\gamma \varepsilon G(\hat{\mathbf{T}}_\varepsilon^{x-y}) \leq t) dP(y) \rightarrow \int_{\mathbb{R}} \mathbb{P}(\hat{\Omega}_{x-y}^*) dP(y) \cdot \frac{t}{1+t}$$

by the  $P = \delta_0$  case of Theorem 2 and dominated convergence and analogously for Theorem 3.

Therefore, for the remainder of this section we assume that  $S'_0 = 0$ .

Next, we observe that our key lemma (Lemma 1) can be adapted as follows. Let  $\Gamma_{R,K}^x$  be the event defined by

$$\Gamma_{R,K}^x := \{\forall i = 1, \dots, K : S_i \neq x \text{ and } |S_i| \leq R\}.$$

**LEMMA 8.** *Let  $c > 0$ , and let, for every  $\varepsilon > 0$ ,  $n_\varepsilon$  be the integer such that  $G(n_\varepsilon) \leq \frac{c}{\varepsilon} < G(n_\varepsilon + 1)$ .*

*For arbitrary  $\gamma'$  and  $\gamma''$  such that  $0 < \gamma' < 2f_X(0) < \gamma''$  we have*

$$\mathbb{P}(\Gamma_{R,K}^x) \geq \mathbb{P}(\mathbf{T}_\varepsilon^x > n_\varepsilon) + \mathbb{P}(\Gamma_{R,K}^x) \gamma' \varepsilon \sum_{k=m_\varepsilon}^{n_\varepsilon} \frac{\mathbb{P}(\tau_\varepsilon > n_\varepsilon - k)}{A_k} + o_{\varepsilon,R,K}(1)$$

and

$$\mathbb{P}(\Gamma_{R,K}^x) \leq \mathbb{P}(\mathbf{T}_\varepsilon^x > n_\varepsilon) + \mathbb{P}(\Gamma_{R,K}^x) \gamma'' \varepsilon \sum_{k=m_\varepsilon}^{n_\varepsilon} \frac{\mathbb{P}(\tau_\varepsilon > n_\varepsilon - k)}{A_k} + o_{\varepsilon,R,K}(1).$$

**PROOF.** We have the following analog of formula (8):

$$(17) \quad \mathbb{P}(\Gamma_{R,K}^x) = \sum_{k=0}^{n_\varepsilon} p_k^{x,-} = \sum_{k=0}^{n_\varepsilon} p_k^{x,+}$$

with

$$p_0^{x,\pm} := \mathbb{P}(\Gamma_{R,K}^x \cap \{\forall \ell = 1, \dots, n_\varepsilon : |S_\ell - x| \geq \varepsilon \pm 2\varepsilon^2\})$$

and

$$p_k^{x,\pm} := \mathbb{P}(\Gamma_{R,K}^x \cap \{|S_k - x| < \varepsilon \pm 2\varepsilon^2 \text{ and } \forall \ell = k+1, \dots, n_\varepsilon : |S_\ell - x| \geq \varepsilon \pm 2\varepsilon^2\}).$$

We follow the proof of Lemma 1.

(i) Observe first that

$$p_0^{x,-} \geq \mathbb{P}(\Gamma_{R,K}^x \cap \{\mathbf{T}_\varepsilon^x > n_\varepsilon\}).$$

Now consider indices with  $m_\varepsilon \leq k \leq n_\varepsilon$ . With the same set  $\mathcal{A}$  as in the proof of Lemma 1, we find, arguing as in (9), that

$$\begin{aligned} p_k^{x,-} &\geq \sum_{a \in \mathcal{A}} \mathbb{P}(\Gamma_{R,K}^x \cap \{S_k - x \in Q_a \text{ and } \forall \ell = k+1, \dots, n_\varepsilon : |S_\ell - x| \geq \varepsilon - 2\nu\}) \\ &\geq \sum_{a \in \mathcal{A}} \mathbb{P}(\Gamma_{R,K}^x \cap \{S_k - x \in Q_a\}) \mathbb{P}(\forall \ell = 1, \dots, n_\varepsilon - k : |S_\ell + a| \geq \varepsilon - \nu). \end{aligned}$$

A proof parallel to that of (10) shows that

$$\mathbb{P}(\Gamma_{R,K}^x \cap \{S_k - x \in Q_a\}) \geq \mathbb{P}(\Gamma_{R,K}^x) \frac{\gamma' \nu}{A_k},$$

if  $\varepsilon$  is sufficiently small. Therefore,

$$\begin{aligned} p_k^{x,-} &\geq \mathbb{P}(\Gamma_{R,K}^x) \frac{\gamma' \nu}{A_k} \sum_{a \in \mathcal{A}} \mathbb{P}(\forall \ell = 1, \dots, n_\varepsilon - k : |S_\ell + a| \geq \varepsilon - \nu) \\ &\geq \mathbb{P}(\Gamma_{R,K}^x) \frac{\gamma' \varepsilon}{A_k} (\mathbb{P}(\tau_\varepsilon > n_\varepsilon - k) - 8\nu), \end{aligned}$$

where the second step uses an estimate contained in (11). Continuing as in the proof of Lemma 1, we obtain the first assertion of our lemma.

(ii) Similar adaptations give the second assertion of the lemma.  $\square$

We can now complete the proofs of our main distributional limit theorems:

**PROOF OF THEOREM 2.** We go back to Lemmas 2 and 3, observing that we already have (6) at our disposal. Take  $t \in (0, \infty)$ ,  $R, K \geq 1$  and  $\gamma' < 2f_X(0) < \gamma''$ . For  $\varepsilon > 0$  let  $m_\varepsilon := (\log \varepsilon)^4$  and choose  $n_\varepsilon$ , such that  $G(n_\varepsilon) \leq \frac{t}{\gamma \varepsilon} \leq G(n_\varepsilon + 1)$ , meaning that  $\mathbb{P}(\varepsilon \gamma G(\mathbf{T}_\varepsilon^x) > t) \sim \mathbb{P}(\mathbf{T}_\varepsilon^x > n_\varepsilon)$ .

In view of (6), the estimate (12) of Lemma 2 becomes

$$\liminf_{\varepsilon \rightarrow 0} \varepsilon \sum_{k=m_\varepsilon}^{n_\varepsilon} \frac{\mathbb{P}(\tau_\varepsilon > n_\varepsilon - k)}{A_k} \geq \frac{\mathbb{P}(\Omega^*)}{\gamma} \frac{t}{1+t}.$$



Combining this with the first part of Lemma 8 leads to

$$\limsup_{\varepsilon \rightarrow 0} \mathbb{P}(\mathbf{T}_\varepsilon^x > n_\varepsilon) \leq \mathbb{P}(\Gamma_{R,K}^*) \left( 1 - \frac{\gamma'}{2f_X(0)} \frac{t}{1+t} \right) + o_{R,K}(1).$$

Successively letting  $R \rightarrow \infty$ , then  $K \rightarrow \infty$  and finally  $\gamma' \rightarrow 2f_X(0)$ , we obtain

$$\limsup_{\varepsilon \rightarrow 0} \mathbb{P}(\mathbf{T}_\varepsilon^x > n_\varepsilon) \leq \frac{\mathbb{P}(\Omega_x^*)}{1+t}.$$

To get the corresponding lower bound, parallel to (13), we have

$$\begin{aligned} \mathbb{P}(\mathbf{T}_\varepsilon^x > 2n_\varepsilon) + \mathbb{P}(\Gamma_{R,K}^x) \gamma'' \varepsilon \sum_{k=m_\varepsilon}^{2n_\varepsilon} \frac{\mathbb{P}(\tau_\varepsilon > 2n_\varepsilon - k)}{A_k} \\ \leq \mathbb{P}(\mathbf{T}_\varepsilon^x > n_\varepsilon) + t \frac{\gamma''}{\gamma} \mathbb{P}(\Gamma_{R,K}^x) \mathbb{P}(\tau_\varepsilon > n_\varepsilon) + o(1). \end{aligned}$$

Together with the second part of Lemma 8 (with  $n_\varepsilon$  replaced by  $2n_\varepsilon$ ) and (6), this implies

$$\liminf_{\varepsilon \rightarrow 0} \mathbb{P}(\mathbf{T}_\varepsilon^x > n_\varepsilon) \geq \frac{\mathbb{P}(\Omega_x^*)}{1+t}$$

completing the proof.  $\square$

**PROOF OF THEOREM 3.** We fix  $t \in (0, \infty)$ , and choose  $n_\varepsilon$  such that  $G(n_\varepsilon) \leq \frac{\beta t}{\gamma \varepsilon} < G(n_\varepsilon + 1)$ .

According to the proof of Theorem 5 [see, in particular, (15) in Lemma 5], we know that for  $m_\varepsilon$  with  $m_\varepsilon = o(n_\varepsilon)$ ,

$$\lim_{\varepsilon \rightarrow 0} \varepsilon \sum_{k=m_\varepsilon}^{n_\varepsilon} \frac{\mathbb{P}(\tau_\varepsilon > n_\varepsilon - k)}{A_k} = \frac{\mathbb{P}(\Omega^*)}{\gamma} \Pr(Y \geq t) =: \psi,$$

where  $Y = \Gamma(\frac{1}{\beta})^{-1} \mathcal{E} \mathcal{G}_{1/\beta}^{1/\beta}$  is the limiting random variable of the  $\gamma \beta^{-1} \varepsilon G(\tau_\varepsilon)$ . Therefore, Lemma 8 implies that for  $R, K \geq 1$  and  $\gamma' < 2f_X(0) < \gamma''$ ,

$$\limsup_{\varepsilon \rightarrow 0} \mathbb{P}(\mathbf{T}_\varepsilon^x > n_\varepsilon) \leq \mathbb{P}(\Gamma_{R,K}^x) (1 - \gamma' \psi) + o_{R,K}(1)$$

and

$$\liminf_{\varepsilon \rightarrow 0} \mathbb{P}(\mathbf{T}_\varepsilon^x > n_\varepsilon) \geq \mathbb{P}(\Gamma_{R,K}^x) (1 - \gamma'' \psi) + o_{R,K}(1).$$

Since  $\lim_{K \rightarrow +\infty} \lim_{R \rightarrow +\infty} \mathbb{P}(\Gamma_{R,K}^x) = \mathbb{P}(\Omega_x^*)$ , we get

$$\mathbb{P}(\Omega_x^*) (1 - \gamma'' \psi) \leq \liminf_{\varepsilon \rightarrow 0} \mathbb{P}(\mathbf{T}_\varepsilon^x > n_\varepsilon) \leq \limsup_{\varepsilon \rightarrow 0} \mathbb{P}(\mathbf{T}_\varepsilon^x > n_\varepsilon) \leq \mathbb{P}(\Omega_x^*) (1 - \gamma' \psi)$$

and hence

$$\lim_{\varepsilon \rightarrow 0} \mathbb{P}(\mathbf{T}_\varepsilon^x > n_\varepsilon) = \mathbb{P}(\Omega_x^*)(1 - 2f_X(0)\psi) = \mathbb{P}(\Omega_x^*) \Pr(Y > t)$$

as required.  $\square$

PROOF OF COROLLARY 1. This is an  $\alpha = 2$  case with  $A_n = \sqrt{n}$  and  $f_X(0) = \frac{1}{\sqrt{2\pi}}$ . Recalling that  $\mathcal{G}_{1/2} = \frac{1}{2N^2}$  in distribution (cf. Example XIII.3.b of [7]) proves our claim.  $\square$

## REFERENCES

- [1] BINGHAM, N. H., GOLDIE, C. M. and TEUGELS, J. L. (1987). *Regular Variation. Encyclopedia of Mathematics and Its Applications* **27**. Cambridge Univ. Press, Cambridge. [MR0898871](#)
- [2] BRESSAUD, X. and ZWEIMÜLLER, R. (2001). Non exponential law of entrance times in asymptotically rare events for intermittent maps with infinite invariant measure. *Ann. Henri Poincaré* **2** 501–512. [MR1846853](#)
- [3] BRETAGNOLLE, J. and DACUNHA-CASTELLE, D. (1968). Théorèmes limites à distance finie pour les marches aléatoires. *Ann. Inst. H. Poincaré Sect. B (N.S.)* **4** 25–73. [MR0235598](#)
- [4] CHELIOTIS, D. (2006). A note on recurrent random walks. *Statist. Probab. Lett.* **76** 1025–1031. [MR2269338](#)
- [5] CHUNG, K.-L. and ERDÖS, P. (1947). On the lower limit of sums of independent random variables. *Ann. of Math. (2)* **48** 1003–1013. [MR0023010](#)
- [6] DVORETZKY, A. and ERDÖS, P. (1951). Some problems on random walk in space. In *Proceedings of the Second Berkeley Symposium on Mathematical Statistics and Probability, 1950* 353–367. Univ. California Press, Berkeley. [MR0047272](#)
- [7] FELLER, W. (1971). *An Introduction to Probability Theory and Its Applications. Vol. II*, 2nd ed. Wiley, New York. [MR0270403](#)
- [8] PÈNE, F. and SAUSSOL, B. (2009). Quantitative recurrence in two-dimensional extended processes. *Ann. Inst. Henri Poincaré Probab. Stat.* **45** 1065–1084. [MR2572164](#)
- [9] STONE, C. (1965). A local limit theorem for nonlattice multi-dimensional distribution functions. *Ann. Math. Statist.* **36** 546–551. [MR0175166](#)

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